Anytime-valid testing and confidence intervals in contingency tables and beyond Turner and Peter Grü —Rosanne . A/B Testing Worksop 2022

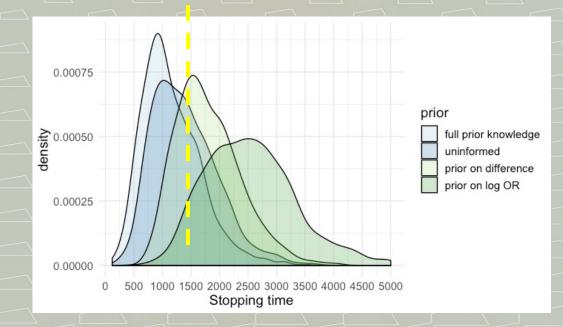
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Goal: tests that can be used under optional stopping (sequential research), with a notion of effect size data collection \rightarrow 500

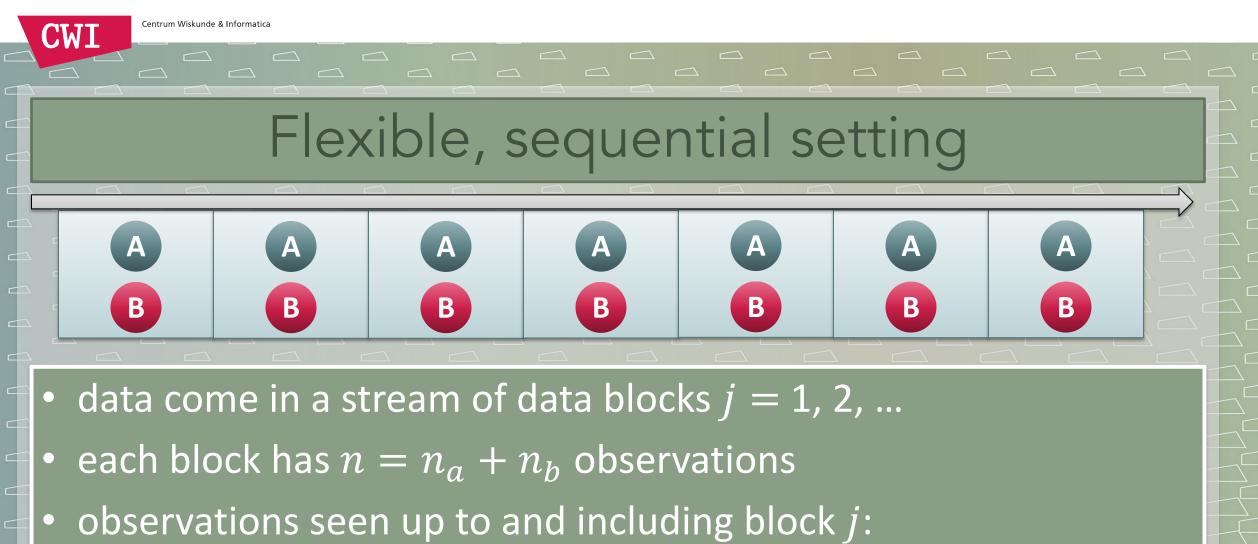
Example: SWEPIS study on stillbirth

- Comparing perinatal death in labour induction at 41 or 42 weeks
- Stopped after ±1380 births in each group: 6 perinatal deaths in 42 weeks group
- Sequential test with balanced design: would often have stopped earlier

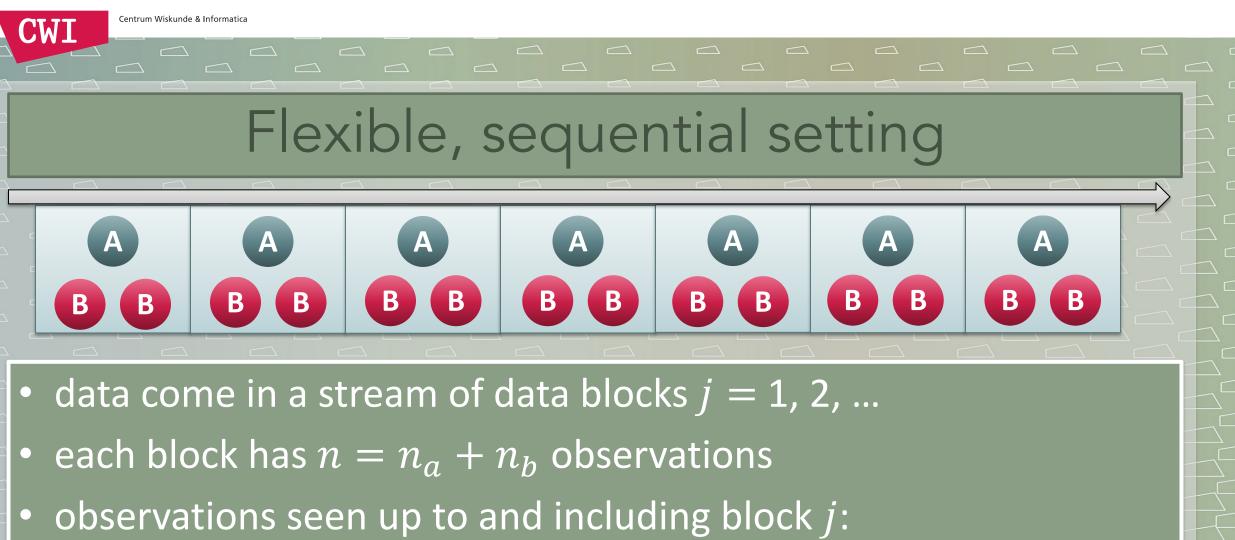
Simulated stopping times with and without using knowledge from previous studies in sequential test*



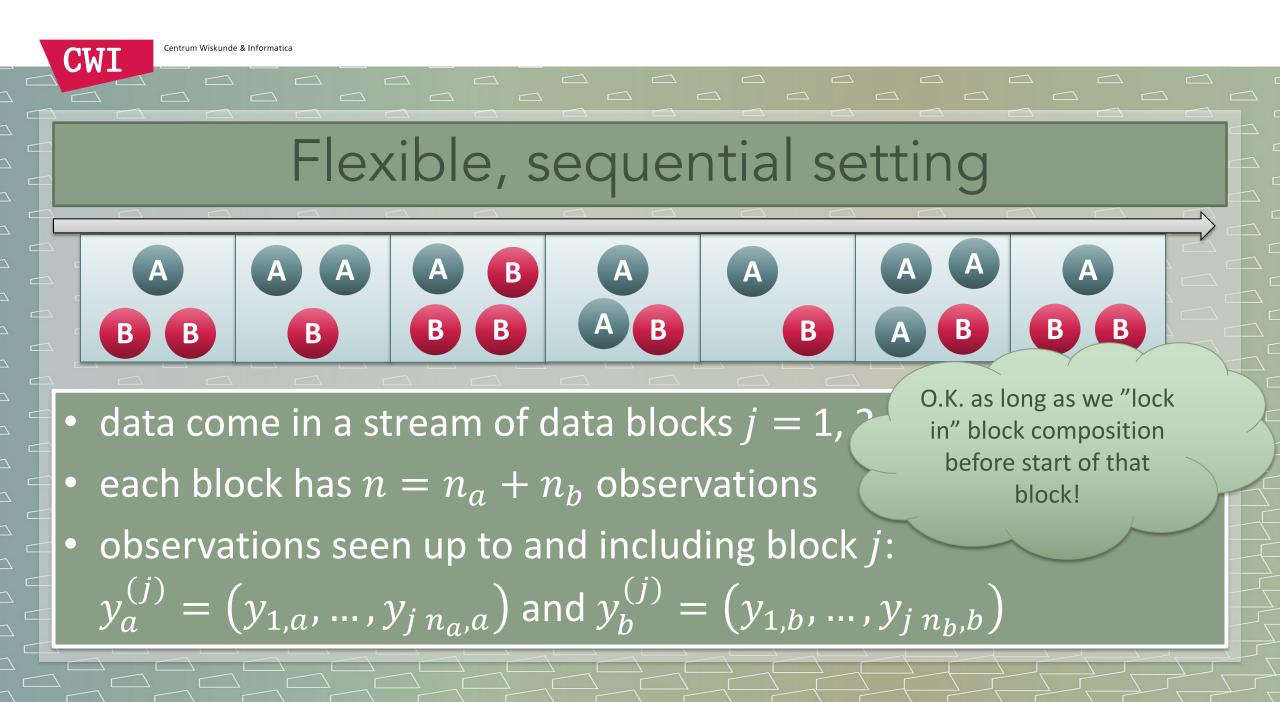
SWEPIS study: Wennerholm et al. published in bmj, 367, 2019. Figure: adapted from Turner et al., 2021



$$y_a^{(j)} = (y_{1,a}, \dots, y_{j n_a, a}) \text{ and } y_b^{(j)} = (y_{1,b}, \dots, y_{j n_b, b})$$



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Running example: 2x2 contingency table setting

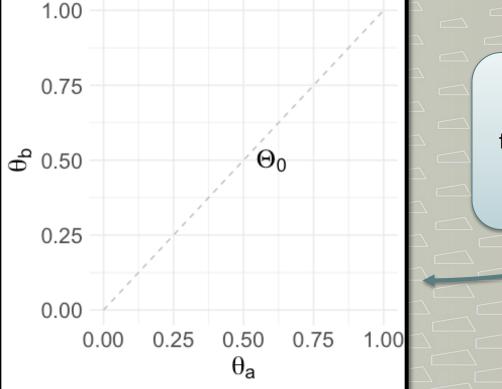
		Strategy		
2x2 contingency table		А	В	
Outcome	Success	S(A)	S(B)	
Outc	Failure	F(A)	F(B)	

Do success probabilities differ between strategies?

- \mathcal{H}_0 : observations $Y \in \{0,1\}$ independent of strategy $X \in \{a, b\}$
- Equivalently, when $Y_x \stackrel{i.i.d.}{\sim}$ Bernoulli(θ_x): $\mathcal{H}_0: \theta_a = \theta_b.$

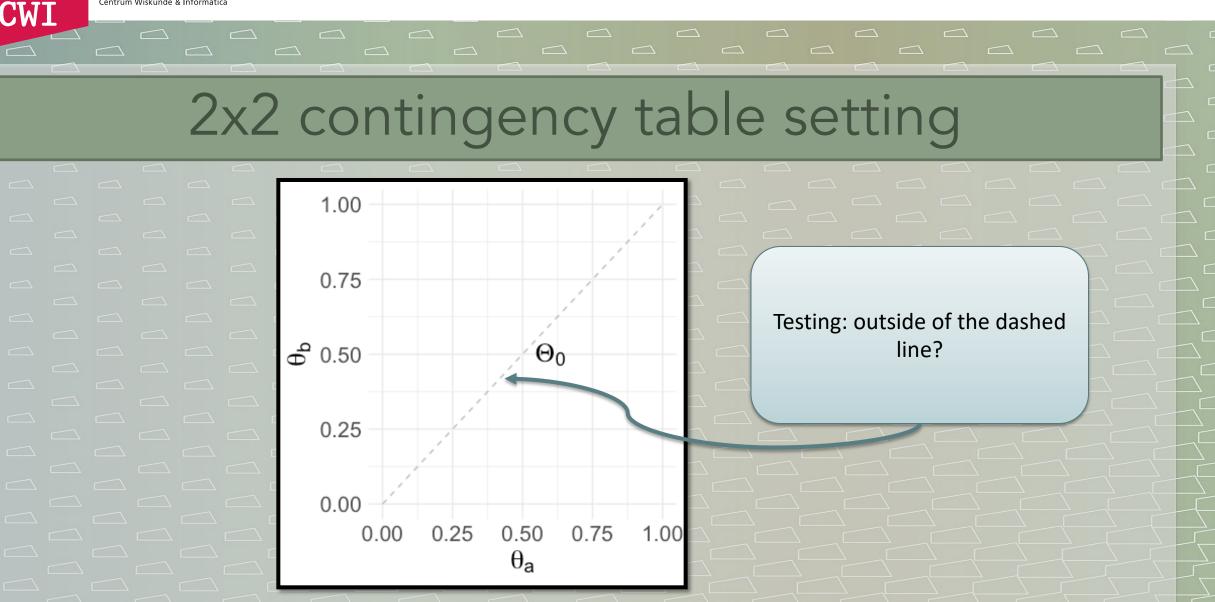
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2x2 contingency table setting



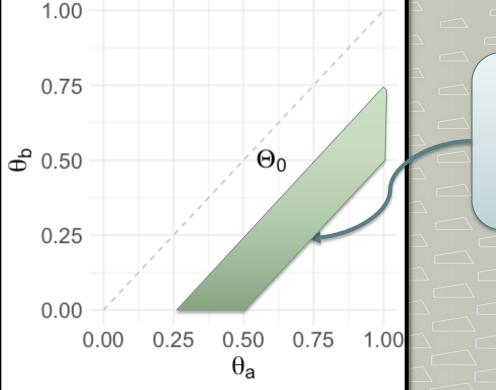
"True" success probabilities for each strategy somewhere in the unit square





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Estimating: somewhere in the shaded area?

Tool for analyzing sequential data: E-variables*

- Nonnegative RV *S*, where for all $P_0 \in \mathcal{H}_0$: $\mathbb{E}_{P_0}[S] \leq 1$
- Straightforward implementation in test: reject \mathcal{H}_0 iff $S \ge \alpha^{-1}$
- Type-I error guarantee at α (e.g. $\alpha = 0.05$, reject if $S \ge 20$)

Betting interpretation \mathcal{H}_0 true? Expect no profit



High profit? Reject $\mathcal{H}_{\mathbf{0}}$



Point alternative 2 data streams: nice general expression!

Point $\mathcal{H}_1 P_{\theta_a, \theta_b}$ (Turner, 2021):

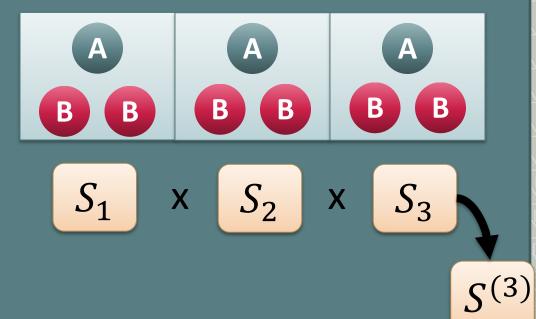
$$S(Y^{(1)}) \coloneqq \prod_{i=1}^{n_a} \frac{p_{\theta_a}(Y_{i,a})}{p_{\theta_0}(Y_{i,a})} \prod_{i=1}^{n_b} \frac{p_{\theta_b}(Y_{i,b})}{p_{\theta_0}(Y_{i,b})}$$

E-variable when we choose $\theta_0 = (n_a/n)\theta_a + (n_b/n)\theta_b$

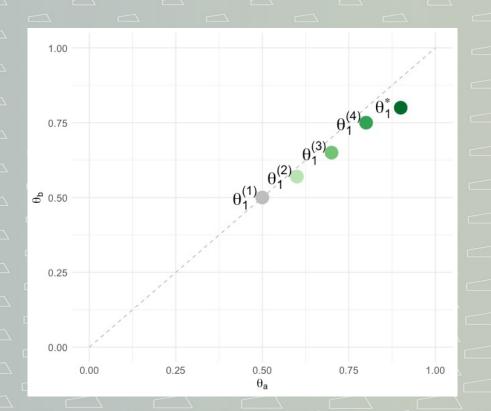
E-process for two data streams

- Can make an e-process: multiply Evalues for all data blocks m
 - $S^{(m)}(Y^{(m)}) \coloneqq \prod_{j=1}^{j} S(Y_j)$
- For arbitrary stopping rule (E-value \geq 20, no money for further experiment, etc..): $P_0(\exists m: S^{(m)}(Y^{(m)}) \geq \alpha^{-1}) \leq \alpha$

Key: multiplying E-values yields another E-value



Learn parameter for \mathcal{H}_1



- Can learn estimate $(\hat{\theta}_a, \hat{\theta}_b)$ of true alternative before each new data block, based on past data
 - Maximum likelihood
 - MAP estimator
 - Posterior mean, ...
- Restrict search space based on expert knowledge

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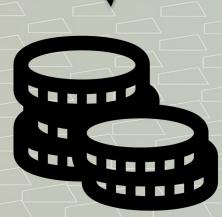
Learn parameter for \mathcal{H}_1



- Can learn estimate $(\hat{\theta}_a, \hat{\theta}_b)$ of true alternative before each new data block, based on past data
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Evidence against \mathcal{H}_1 and Type-II error

- GRO criterion: in sequential experiments: optimize "growth rate" of E-variable, E_{P1} [log S] (Grünwald, 2019)
- Minimize notion of regret: loss of capital growth under alternative due to not knowing true P₁.
- Closely connected to optimizing power



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2x2 E-values vs classical counterpart

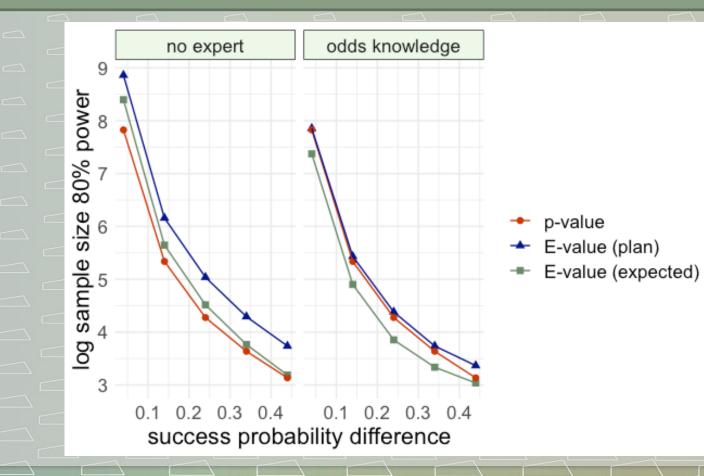


Figure adapted from Turner et al., 2021, figure 4

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2x2 E-values vs classical counterpart

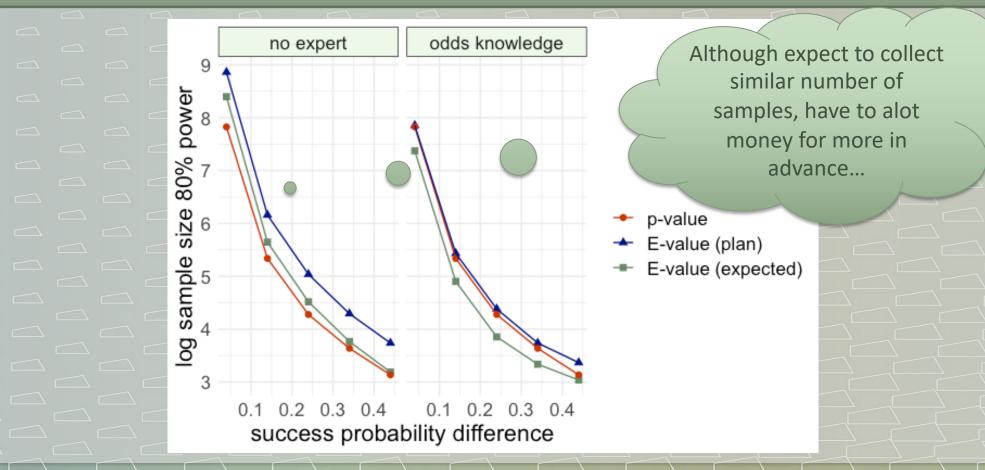


Figure adapted from Turner et al., 2021, figure 4

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2x2 E-values vs classical counterpart

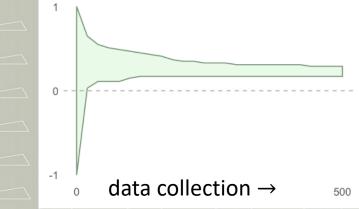
odds knowledge no expert 9 power 8 80% **a** 6 p-value E-value (plan) E-value (expected) On plus side: allowed to continue experiment/ combine with new experiment even years after first experiment has ended! 0.3 0.4 0.1 0.2 0.3 0.4 success probability difference

Figure adapted from Turner et al., 2021, figure 4



Anytime-valid confidence sequences

Update effect size estimate each time a new batch of data has come in, **with coverage guarantee** (real value is in my estimate with some minimum probability)



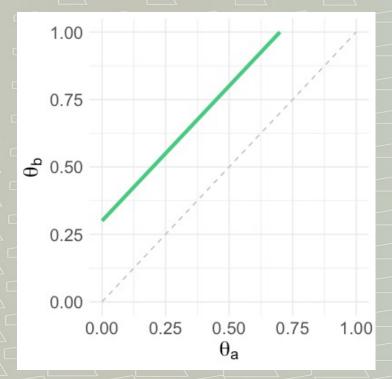
Formally; confidence sequence CS with coverage at level $(1 - \alpha)$: $-P_{\theta_a,\theta_b}(\text{ for any } m = 1, 2, ... : \delta(\theta_a, \theta_b) \notin CS_{(m)}) \leq \alpha$ $-\delta(\theta_a, \theta_b)$: measure of *effect size*

Key: use E-process to test effect size values

- Let $S_{\Theta_0(\delta)}^{(m)}$ be an E-process for testing: $\mathcal{H}_0 \coloneqq \{P_{\theta_0} : \theta_0 \in \Theta_0(\delta)\}$
- Probability of falsely rejecting \mathcal{H}_0 bounded by α (because it is an E-process)!
- Construct anytime-valid confidence sequence $CS_{\alpha,(m)} = \left\{ \delta: S_{\Theta_{\alpha}(\delta)}^{(m)} \leq \frac{1}{\alpha} \right\}$
- \rightarrow gives us the desired coverage at level (1α) .

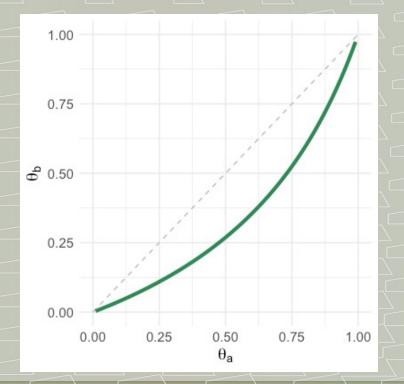
Extension to \mathcal{H}_0 beyond $\theta_a = \theta_b$: examples

Effect size $\delta: (\theta_a, \theta_b) \rightarrow \gamma; \gamma \in \Gamma$. - E.g. Risk Difference: $\delta(\theta_a, \theta_b) = \theta_b - \theta_a, \Gamma = [-1, 1]$ - E.g. Odds Ratio: $\delta(\theta_a, \theta_b) = \frac{\theta_b}{1 - \theta_b} \frac{1 - \theta_a}{\theta_a}, \Gamma = \mathbb{R}^+$ $\Theta_0(\delta) = \{(\theta_a, \theta_b): \theta_b - \theta_a = 0.3\}$



Extension to \mathcal{H}_0 beyond $\theta_a = \theta_b$: examples

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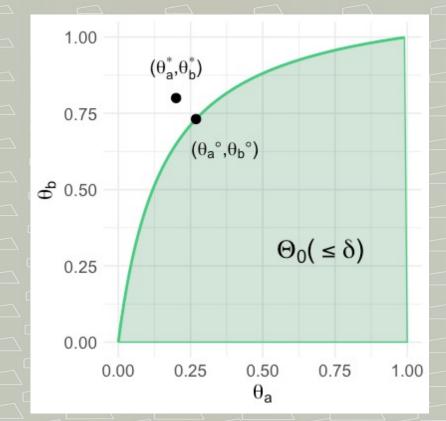


Extension of E-variable for streams to general null hypothesis $\Theta_0(\delta)$ for 2x2 tables

 $S_{\Theta_{0}}(Y^{(1)}) \coloneqq \prod_{i=1}^{n_{a}} \frac{p_{\widehat{\theta}_{a}}(Y_{i,a})}{p_{\theta_{a}^{\circ}}(Y_{i,a})} \prod_{i=1}^{n_{b}} \frac{p_{\widehat{\theta}_{b}}(Y_{i,b})}{p_{\theta_{b}^{\circ}}(Y_{i,b})},$ where $(\theta_{a}^{\circ}, \theta_{b}^{\circ})$ achieve $\min_{(\theta_{a}, \theta_{b}) \in \Theta_{0}(\delta)} D(P_{\widehat{\theta}_{a}, \widehat{\theta}_{b}}(Y_{a}^{n_{a}}, Y_{b}^{n_{b}})|P_{\theta_{a}^{\circ}, \theta_{b}^{\circ}}(Y_{a}^{n_{a}}, Y_{b}^{n_{b}}))$ and we estimate the point $(\widehat{\theta}_{a}, \widehat{\theta}_{b})$ as before (Turner, 2022)

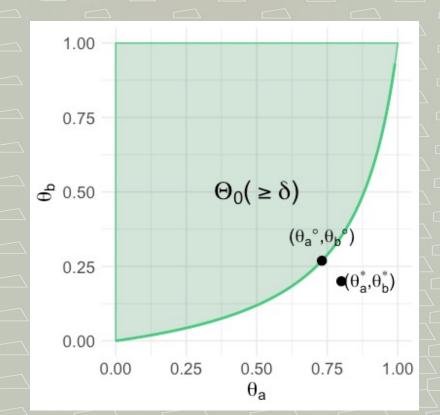
Tricky case: odds ratio and convexity of \mathcal{H}_0

- Need convexity of $\Theta_0(\delta)$ to construct E-variable
- δ > 0 → can estimate lower bound (see figure)
- δ < 0 → can estimate
 upper bound



Tricky case: odds ratio and convexity of \mathcal{H}_0

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 upper bound (see figure)



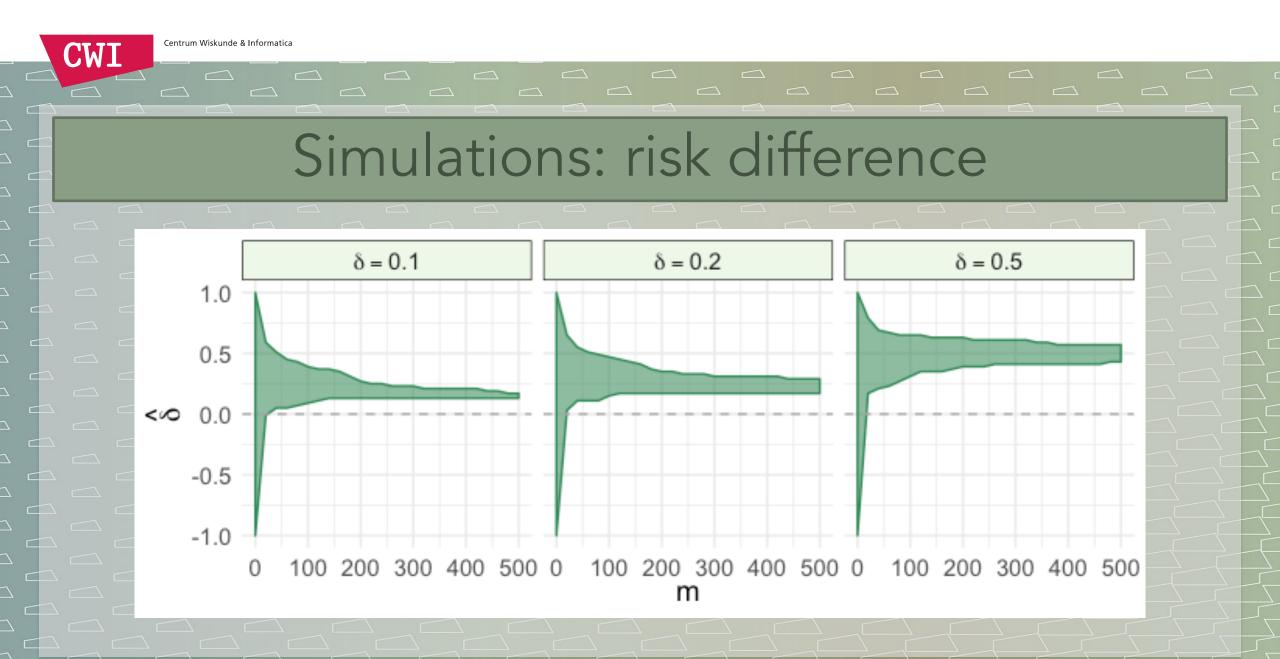


Figure adapted from Turner et al., 2022



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Simulations: risk difference

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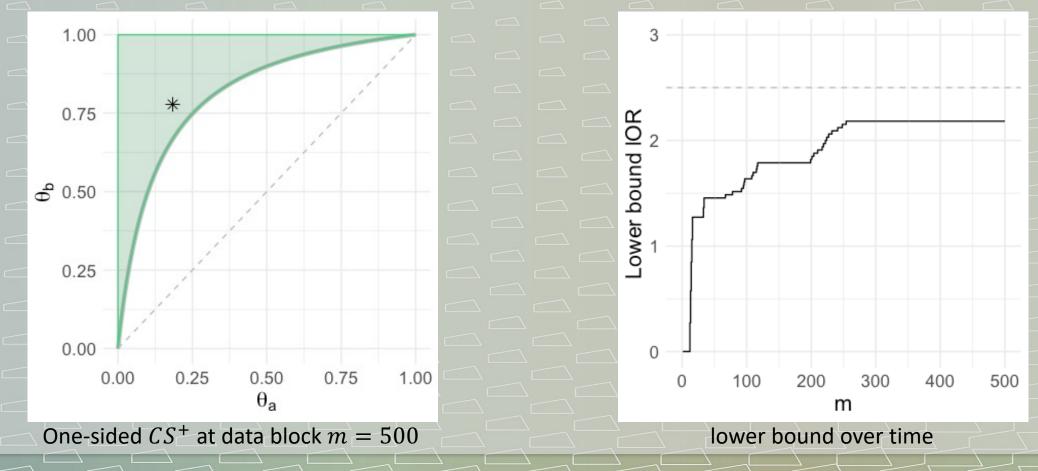


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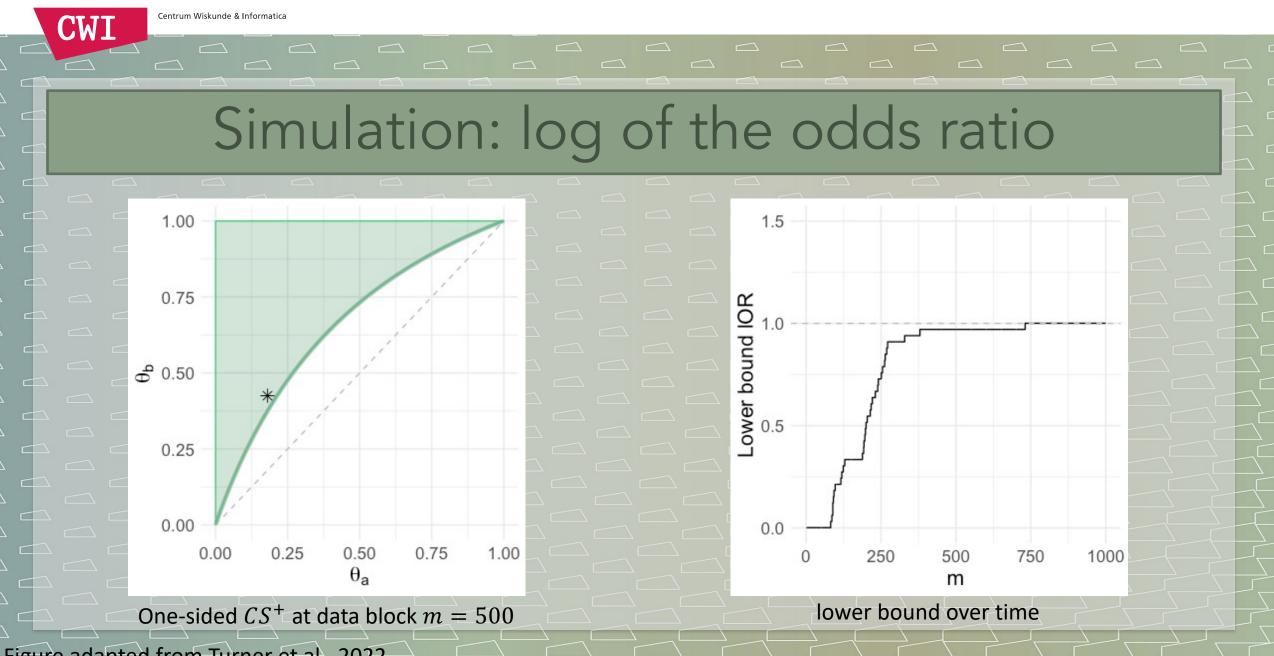


Figure adapted from Turner et al., 2022

Conclusion and novelty

- To our knowledge, really new:
 - flexibility (block size, user-specified notions of effect size)
 - growth rate optimality: expect evidence for H1 to grow as fast as possible during data collection
- Wald's sequential probability ratio test:
 - Probability ratios can be interpreted as "alternative" E-variables
 - Not growth-rate optimal
 - Only allow for testing odds ratio effect size

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Extensions

- Beyond Bernoulli: GRO property? (work by Y. Hao and others)
 Stratified data and conditional independence
 - Use case at UMC Utrecht: real-time psychiatry research and recommendations

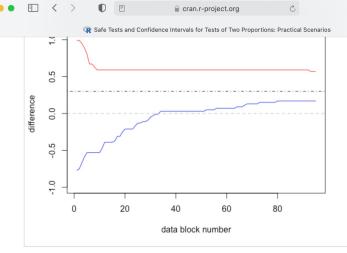
		Strategy	
		А	В
um 1	Success	S(A1)	S(B1)
Stratum	Failure	F(A1)	F(B1)
um 2	Success	S(A2)	S(B2)
Stratum	Failure	F(A2)	F(B2)
nm 3	Success	S(A3)	S(B3)
Stratum	Failure	F(A3)	F(B3)

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R Package and Vignettes

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The above plot shows that after finishing our experiment, 0 is not included in the confidence interval (grey dashed line). The true value, 0.3, remains included. The precision indicates how many difference values between -1 and 1 are checked while building the confidence sequence. It is recommended to set this value to 100 (default).

The code below can be used to check that our confidence sequence indeed offers the $1 - \alpha$ guarantee and includes the difference between the two success probabilities of 0.3 in at least 95% of simulated scenarios:

print(coverageSimResult)
#> [1] 0.974

- In R console: install.packages (
 - "safestats")
 - <u>https://CRAN.R-</u> project.org/package=safesta

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Further reading and references

- On the theory of E-values:
 - P.D. Grünwald, R. de Heide and W. Koolen (2019) on ArXiv:
 - V. Vovk and R. Wang (2021). E-values: Calibration, combination, and applications. Annals of Statistics.
 - G. Shafer (2021). Testing by betting: A strategy for statistical and scientific communication. Journal of the Royal Statistical Society, Series A.
- On implementations of E-values:
 - R.J. Turner, A. Ly and P.D. Grünwald (2021) on ArXiv:2106.02693
 - R.J. Turner and P.D. Grünwald (2022) on ArXiv:2203.09785
 - R software: <u>https://CRAN.R-project.org/package=safestats</u>